Generalized Linear Mixed Models

Analysis of Ecological and Environmental Data

QERM 514

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Goals for today

- Understand the structural components of generalized linear mixed models
- Understand the options for fitting GLMMs and their pros and cons
- Understand some of the diagnostics available for evaluating GLMM fits

Forms of linear models



GLMMs combine the flexibility of non-normal distributions (GLMs) with the ability to address correlations among observations and nested data structures (LMMs)

Good news

 $\cdot\,$ these extensions follow similar methods to GLMs and LMMs $\,$

Bad news

- \cdot these models are on the frontier of statistical research
- existing documentation is rather technical
- multiple approaches for fitting models; some with different results

Just like GLMs, GLMMs have three components:

- 1. Distribution of the data $f(y; \theta)$
- 2. Link function $g(\eta)$
- 3. Linear predictor η

Linear predictor for a GLM

We can write the linear predictor for GLMs as

where the β_i are fixed effects of the covariates x_i

Linear predictor for a GLMM

For GLMMs, our linear predictor also includes random effects

$$\eta = \beta_0 + \beta_1 x_1 + \dots + \beta_k x_k + \alpha_0 + \alpha_1 z_1 + \dots + \alpha_l z_l$$
$$\Downarrow$$
$$\eta = \mathbf{X}\boldsymbol{\beta} + \mathbf{Z}\boldsymbol{\alpha}$$

where the β_i are fixed effects of the covariates x_i

Survival of fish $s_{i,j}$ as a function of length $x_{i,j}$ in some location j

data distribution: $y_{i,j} \sim \text{Binomial}(N_{i,j}, s_{i,j})$ link function: $\text{logit}(s_{i,j}) = \log\left(\frac{s_{i,j}}{1 - s_{i,j}}\right) = \mu_{i,j}$ linear model: $\mu_{i,j} = (\beta_0 + \alpha_j) + \beta_1 x_{i,j}$ $\alpha_j \sim N(0, \sigma_{\delta}^2)$

Best practices suggest we try to keep things simple

Why? Because GLMMs involve solving an integral with no analytical solution

Likelihood for GLMMs

Recall that we think of likelihoods in terms of the observed data

But the random effects in our model are *unobserved* random variables, so we need to integrate them out of the likelihood

Likelihood for GLMMs

The likelihood for a GLMM involves integrating over all possible random effects

$$\mathcal{L}(y; \boldsymbol{\beta}, \boldsymbol{\phi}, \boldsymbol{\nu}) = \prod_{i} \int \underbrace{f_d(y; \boldsymbol{\beta}, \boldsymbol{\phi}, \boldsymbol{\alpha})}_{\text{distn for data}} \underbrace{f_r(\boldsymbol{\alpha}; \boldsymbol{\nu})}_{\text{distn for RE}} d\boldsymbol{\alpha}$$

If $f(y; \beta, \phi, \alpha)$ is not Gaussian, we cannot remove it from the likelihood, which makes it *very* difficult to compute

Approaches to fitting GLMMs

To avoid the integral, we will consider 3 methods that approximate the likelihood

They all have pros and cons so it's not possible to pick the "best"

Penalized quasi-likelihood (PQL) uses a Taylor series expansion to approximate the linear predictor as an LMM

$$g(\boldsymbol{\mu}) = \boldsymbol{\eta}$$

= $\mathbf{X}\boldsymbol{\beta} + \mathbf{Z}\boldsymbol{\alpha}$
 \Downarrow
 $g(\mathbf{y}) \approx g(\boldsymbol{\mu}) + g'(\boldsymbol{\mu})(\mathbf{y} - \boldsymbol{\mu})$
 $\approx \mathbf{X}\boldsymbol{\beta} + \mathbf{Z}\boldsymbol{\alpha} + g'(\boldsymbol{\mu})\boldsymbol{\epsilon}$

The conditional variance of the data in a GLMM is then

$$g(\mathbf{y}) \approx \mathbf{X}\boldsymbol{\beta} + \mathbf{Z}\boldsymbol{\alpha} + g'(\boldsymbol{\mu})\boldsymbol{\epsilon}$$
$$\Downarrow$$
$$g(\mathbf{y}) - \mathbf{X}\boldsymbol{\beta} \approx \mathbf{Z}\boldsymbol{\alpha} + \boldsymbol{\epsilon}g'(\boldsymbol{\mu})$$
$$\Downarrow$$
$$\forall$$
$$Var\left(g(\mathbf{y}) - \mathbf{X}\boldsymbol{\beta}\right) \approx Var\left(\mathbf{Z}\boldsymbol{\alpha}\right) + Var\left(\boldsymbol{\epsilon}g'(\boldsymbol{\mu})\right)$$

The conditional variance of the data in a GLMM is then

$$g(\mathbf{y}) \approx \mathbf{X}\boldsymbol{\beta} + \mathbf{Z}\boldsymbol{\alpha} + g'(\boldsymbol{\mu})\boldsymbol{\epsilon}$$
$$\Downarrow$$
$$g(\mathbf{y}) - \mathbf{X}\boldsymbol{\beta} \approx \mathbf{Z}\boldsymbol{\alpha} + \boldsymbol{\epsilon}g'(\boldsymbol{\mu})$$
$$\Downarrow$$
$$\forall$$
$$Var\left(g(\mathbf{y}) - \mathbf{X}\boldsymbol{\beta}\right) \approx Var\left(\mathbf{Z}\boldsymbol{\alpha}\right) + Var\left(\boldsymbol{\epsilon}g'(\boldsymbol{\mu})\right)$$

which is similar to that for an LMM

$$\operatorname{Var}(\mathbf{y} - \mathbf{X}\boldsymbol{\beta}) = \operatorname{Var}(\mathbf{Z}\boldsymbol{\alpha}) + \operatorname{Var}(\boldsymbol{\epsilon})$$

<u>Pros</u>

• fast, flexible, and widely implemented

<u>Cons</u>

- only asymptotically correct
- biased for Binomial and Poisson with small samples
- inference confounded by approximate likelihood

Laplace approximation

Laplace approximation is a long standing (1774) method for computing integrals of the form

 $\int f(x)e^{\lambda g(x)}dx$

This integrand is quite similar to the likelihood of a GLMM based on exponential distributions

Thus, we only need to find the maximum of g(x) and its second derivative, and evaluate them at only one point

Laplace approximation

<u>Pros</u>

- approximation of true likelihood rather than quasi-likelihood
- \cdot more accurate than PQL

<u>Cons</u>

- slower and less flexible than PQL
- \cdot may be impossible to compute for complex models

Gauss-Hermite quadrature

Gauss-Hermite quadrature is an expansion of Laplace approximation where the integrand is evaluated at more than one point

Quadrature is a method for numerically approximating an integral as a weighted sum

$$\int f(u)e^{-u^2}du \approx \sum_i w_i f(u_i)$$

This method works by optimizing the placement and number of the u_i and the choice of the w_i

Gauss-Hermite quadrature

<u>Pros</u>

• More accurate than Laplace

<u>Cons</u>

- Slow and computationally intense
- Limited to a few random effects (one in practice)

Fitting GLMMs

Example

Let's consider a long-term study of invasive brown tree snakes in Guam

Introduced to the island shortly after WWII

Voracious predators on native birds and other vertebrates

Photo by Pavel Kirillov

Our data consist of counts of the number of eggs per female at 23 locations over 14 years

We are interested in the fixed effect of body size and the random effects of location and year

We'll begin with only the effects of body size and location



Penalized quasi-likelihood

We fit PQL models with MASS::glmPQL()

summary(snakes_pql)

```
## Linear mixed-effects model fit by maximum likelihood
## Data: df eggs
   AIC BIC logLik
##
##
   NA NA
                NA
##
## Random effects:
## Formula: ~1 | loc
         (Intercept) Residual
##
## StdDev: 0.5077229 1.183238
##
## Variance function:
## Structure: fixed weights
## Formula: ~invwt
## Fixed effects: eggs ~ size
                 Value Std.Error DF t-value p-value
##
## (Intercept) 1.1247363 0.11687536 210 9.623383
                                                      0
        0.5079533 0.07916825 210 6.416124
## size
                                                      0
## Correlation:
##
       (Intr)
## size -0.069
##
## Standardized Within-Group Residuals:
##
         Min
                     01
                               Med
                                           Q3
                                                     Max
## -1.7744344 -0.7176552 -0.2481373 0.5028263 3.3994803
##
## Number of Observations: 234
## Number of Groups: 23
```

Laplace

We can fit Laplace models with lme4::glmer()

summary(snakes_lap)

```
## Generalized linear mixed model fit by maximum likelihood (Laplace
    Approximation) [glmerMod]
##
## Family: poisson ( log )
## Formula: eggs ~ size + (1 | loc)
     Data: df_eggs
##
##
                   logLik deviance df.resid
##
       ATC
                BIC
##
   1006.7 1017.1 -500.4 1000.7
                                          231
##
## Scaled residuals:
      Min 10 Median 30
##
                                     Max
## -2.1158 -0.8480 -0.2741 0.5931 4.0679
##
## Random effects:
## Groups Name
                  Variance Std.Dev.
## loc (Intercept) 0.2753 0.5247
## Number of obs: 234, groups: loc, 23
##
## Fixed effects:
              Estimate Std. Error z value Pr(>|z|)
##
## (Intercept) 1.09929 0.11726 9.374 < 2e-16 ***
## size
              0.50619 0.06644 7.619 2.56e-14 ***
## ___
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##
       (Intr)
## size -0.054
```

Gauss-Hermite quadrature

We can fit GHQ models with lme4::glmer(..., nAGQ = pts)

```
## fit model
snakes_ghq <- glmer(eggs ~ size + (1 | loc), data = df_eggs,
family = poisson, nAGQ = 20)</pre>
```

summary(snakes_ghq)

Note: this method only works with one random effect

```
## Generalized linear mixed model fit by maximum likelihood (Adaptive
   Gauss-Hermite Quadrature, nAGQ = 20) [glmerMod]
##
## Family: poisson ( log )
## Formula: eggs ~ size + (1 | loc)
##
     Data: df_eggs
##
                   logLik deviance df.resid
##
       ATC
                BIC
##
     397.7 408.1 -195.9
                               391.7
                                          231
##
## Scaled residuals:
      Min 10 Median 30
##
                                     Max
## -2.1159 -0.8479 -0.2739 0.5929 4.0681
##
## Random effects:
## Groups Name
                  Variance Std.Dev.
## loc (Intercept) 0.2761 0.5254
## Number of obs: 234, groups: loc, 23
##
## Fixed effects:
              Estimate Std. Error z value Pr(>|z|)
##
## (Intercept) 1.09919 0.11754 9.352 < 2e-16 ***
## size
              0.50618 0.06681 7.576 3.56e-14 ***
## ___
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##
       (Intr)
## size -0.054
```

Here is a summary of the results from the 3 methods

##		PQL	SE	Laplace	SE	GHQ	SE
##	(Intercept)	1.125	0.117	1.099	0.117	1.099	0.118
##	size	0.508	0.079	0.506	0.066	0.506	0.067
##	location SD	0.508	NA	0.525	NA	0.525	NA

What if we also want to include the random effect of year?

glmmPQL only allows for nested random effects

glmer(..., nAGQ = pts) only allows for one random effect

We can use the Laplace approximation via glmer

Laplace

summary(snakes_lap_2)

```
## Generalized linear mixed model fit by maximum likelihood (Laplace
    Approximation) [glmerMod]
##
## Family: poisson ( log )
## Formula: eqgs ~ size + (1 \mid loc) + (1 \mid year)
     Data: df eggs
##
##
                    logLik deviance df.resid
##
       ATC
                BIC
##
     928.8 942.6 -460.4
                                920.8
                                          230
##
## Scaled residuals:
           10 Median
##
      Min
                              30
                                     Max
## -1.7498 -0.6251 -0.0568 0.5055 3.5431
##
## Random effects:
## Groups Name
                    Variance Std.Dev.
## loc
         (Intercept) 0.2522 0.5022
## year (Intercept) 0.1557 0.3945
## Number of obs: 234, groups: loc, 23; year, 14
##
## Fixed effects:
              Estimate Std. Error z value Pr(>|z|)
##
## (Intercept) 1.03612 0.15518 6.677 2.44e-11 ***
## size
               0.51380 0.07063 7.274 3.48e-13 ***
## ___
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##
       (Intr)
## size -0.048
```



Diagnostics for GLMMs are similar to those for GLMs, but we are limited in our choices

Goodness of fit

Recall our goodness of fit test based on the Pearson's χ^2

$$X^{2} = \sum_{i=1}^{n} \frac{(O_{i} - E_{i})^{2}}{E_{i}} \sim \chi^{2}_{(n-1)}$$

where O_i is the observed count and E_i is the expected count

Pearson's χ^2 statistic

For a binomial distribution

$$X^{2} = \sum_{i=1}^{n} \frac{(y_{i} - n_{i}\hat{p}_{i})^{2}}{n_{i}\hat{p}_{i}}$$

For a Poisson distribution

$$X^{2} = \sum_{i=1}^{n} \frac{(y_{i} - \lambda_{i})^{2}}{\lambda_{i}}$$

Goodness of fit

```
H_0: Our model is correctly specified
```

[1] 0.9986993

The *p*-value is large so we cannot reject H_0

Model diagnostics





For other models, we can calculate the leverages to evaluate potentially extreme values in predictor space

For GLMMs, however, the leverages depend on the estimated variancecovariance matrices of the random effects

Cook's Distance

For other models, we can calculate Cook's distances to identify potentially influential data points

For GLMMs, however, the Cook's distances involve derivatives of the likelihood with respect to the random effects (this is an active area of research)

Inference for fixed effects

We can test the significance of the fixed effects via a χ^2 test by comparing models with and without the effect(s)

```
## Data: df_eggs
## Models:
## snakes_lap_null: eggs ~ (1 | loc) + (1 | year)
## snakes_lap_2: eggs ~ size + (1 | loc) + (1 | year)
## Df AIC BIC logLik deviance Chisq Chi Df Pr(>Chisq)
## snakes_lap_null 3 979.77 990.14 -486.88 973.77
## snakes_lap_2 4 928.81 942.63 -460.40 920.81 52.961 1 3.402e-13
##
## snakes_lap_null
## snakes_lap_2 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Inference for random effects

We can test the significance of the random effects via a χ^2 test by comparing models with and without the effect(s)

```
## Data: df_eggs
## Models:
## snakes_lap_null: eggs ~ size
## snakes_lap_2: eggs ~ size + (1 | loc) + (1 | year)
## Df AIC BIC logLik deviance Chisq Chi Df
## snakes_lap_null 2 1173.92 1180.83 -584.96 1169.92
## snakes_lap_2 4 928.81 942.63 -460.40 920.81 249.11 2
## Pr(>Chisq)
## snakes_lap_null
## snakes_lap_2 < 2.2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1</pre>
```

Overdispersion

As with GLMs, we can check GLMMs for evidence of overdispersion, which we estimate as

$$\hat{c} = \frac{X^2}{n-k}$$

Let's do so for our snake model applied to another data set

Overdispersion

```
## Pearson's X^2 statistic
X2 <- sum((eggs - fitted(snakes_lap))^2 / fitted(snakes_lap))
## number of parameters
k <- length(coef(snakes_lap)) + length(ranef(snakes_lap))
## overdispersion parameter
(c hat <- X2 / (nn - k))</pre>
```

[1] 2.767328

pchisq(deviance(snakes_lap), k, lower.tail = FALSE)

[1] 5.191758e-216

Negative binomial

We can fit neg binomial models using Laplace approximation with lme4::glmer.nb()

summary(snakes_lap_nb)

```
## Generalized linear mixed model fit by maximum likelihood (Laplace
    Approximation) [glmerMod]
##
## Family: Negative Binomial(148955.1) ( log )
## Formula: eqgs ~ size + (1 | loc) + (1 | year)
     Data: df eggs
##
##
                    logLik deviance df.resid
##
       ATC
                BIC
##
     930.8
              948.1 -460.4
                               920.8
                                          229
##
## Scaled residuals:
           10 Median
##
      Min
                               30
                                     Max
## -1.7498 -0.6251 -0.0568 0.5055 3.5431
##
## Random effects:
## Groups Name
                    Variance Std.Dev.
## loc
         (Intercept) 0.2522 0.5022
## year (Intercept) 0.1557 0.3946
## Number of obs: 234, groups: loc, 23; year, 14
##
## Fixed effects:
              Estimate Std. Error z value Pr(>|z|)
##
## (Intercept) 1.03607 0.15537 6.668 2.59e-11 ***
## size
               0.51382 0.07122 7.215 5.41e-13 ***
## ___
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##
       (Intr)
## size -0.048
```

Summary of GLMM methods

Method	Advantages	Disadvantages	R functions
Penalized quasi- likelihood	Flexible, widely implemented	inference may be inappropriate; potentially biased	MASS::glmmPQL
Laplace approximation	More accurate than PQL	Slower and less flexible than PQL	<pre>lme4::glmer glmmsr::glmm glmmML::glmmML</pre>
Gauss-Hermite quadrature	More accurate than Laplace	Slower than Laplace; limited random effects	<pre>lme4::glmer glmmsr::glmm glmmML::glmmML</pre>

Adapted from Bolker et al (2009)