

# Design matrices for models

Analysis of Ecological and Environmental Data

QERM 514

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# Goals for today

- Understand how to create design matrices for use in linear models
- Recognize the different coding schemes for factor models
- See how to use `model.matrix()` for creating & extracting design matrices

# Models in matrix form

Recall the matrix form for our linear models, where

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{e}$$
$$\mathbf{e} \sim \text{MVN}(\mathbf{0}, \boldsymbol{\Sigma})$$

# Models in matrix form

Let's write out this model in more detail

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{e}$$

⇓

$$\begin{bmatrix} \mathbf{y}_1 \\ \mathbf{y}_2 \\ \vdots \\ \mathbf{y}_n \end{bmatrix} = \begin{bmatrix} 1 & \mathbf{x}_{1,1} & \cdots & \mathbf{x}_{n,1} \\ 1 & \mathbf{x}_{1,2} & \cdots & \mathbf{x}_{n,2} \\ \vdots & \vdots & \ddots & \vdots \\ 1 & \mathbf{x}_{1,n} & \cdots & \mathbf{x}_{n,n} \end{bmatrix} \begin{bmatrix} \boldsymbol{\beta}_0 \\ \boldsymbol{\beta}_1 \\ \vdots \\ \boldsymbol{\beta}_n \end{bmatrix} + \begin{bmatrix} \mathbf{e}_1 \\ \mathbf{e}_2 \\ \vdots \\ \mathbf{e}_n \end{bmatrix}$$

The columns in  $\mathbf{X}$  define the *design* of the analysis

# Ordinary least squares

Also recall that we can use  $\mathbf{X}$  to solve for  $\hat{\mathbf{y}}$

$$\begin{aligned}\hat{\mathbf{y}} &= \mathbf{X}\hat{\boldsymbol{\beta}} \\ &= \mathbf{X} \left( (\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{y} \right) \\ &= \underbrace{\mathbf{X}(\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top}_{\mathbf{H}} \mathbf{y} \\ &= \mathbf{H}\mathbf{y}\end{aligned}$$

Understanding the form of  $\mathbf{X}$  is critical to our inference

# A simple starting point

Data = (Deterministic part) + (Stochastic part)

# Types of linear models

We classify linear models by the form of their deterministic part

Discrete predictor → ANalysis Of VAriance (ANOVA)

Continuous predictor → Regression

Both → ANalysis of COVAriance (ANCOVA)

# Possible models for growth of fish

Model	Description
$\text{growth}_i = \beta_0 + \beta_{1,\text{species}} + \epsilon_i$	1-way ANOVA
$\text{growth}_i = \beta_0 + \beta_{1,\text{species}} + \beta_{2,\text{tank}} + \epsilon_i$	2-way ANOVA
$\text{growth}_i = \beta_0 + \beta_1 \text{ration}_i + \epsilon_i$	simple linear regression
$\text{growth}_i = \beta_0 + \beta_1 \text{ration}_i + \beta_2 \text{temperature}_i + \epsilon_i$	multiple regression
$\text{growth}_i = \beta_0 + \beta_{1,\text{species}} + \beta_2 \text{ration}_i + \epsilon_i$	ANCOVA

---



# Defining models with $\mathbf{X}$

Mean only

What would  $\mathbf{X}$  look like for a simple model of the data  $\mathbf{y}$  that included a mean only?

$$\mathbf{y} = \boldsymbol{\mu} + \mathbf{e}$$

# Defining models with $X$

Mean only

Let's start by rewriting our model as

$$\begin{aligned} y &= \beta_0 + \mathbf{e} \\ &= \begin{bmatrix} \beta_0 \\ \beta_0 \\ \vdots \\ \beta_0 \end{bmatrix} + \mathbf{e} \end{aligned}$$

# Defining models with $\mathbf{X}$

Mean only

$$\mathbf{y} = \begin{bmatrix} 1 \\ 1 \\ \vdots \\ 1 \end{bmatrix} \beta_0 + \mathbf{e}$$
$$= \mathbf{X}\boldsymbol{\beta} + \mathbf{e}$$

with  $\mathbf{X} = [1 \ 1 \ \cdots \ 1]^\top$  and  $\boldsymbol{\beta} = [\beta_0]$

# Defining models with $\mathbf{X}$

## Regression

What would  $\mathbf{X}$  look like for a regression model with 2 predictors?

$$y_i = \beta_0 + \beta_1 x_{1,i} + \beta_2 x_{2,i} + e_i$$

⇓?

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{e}$$

# Defining models with $X$

Regression

$$y = X\beta + e$$

$\Downarrow$

$$\begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{bmatrix} = \begin{bmatrix} 1 & \mathbf{x}_{1,1} & \mathbf{x}_{2,1} \\ 1 & \mathbf{x}_{1,2} & \mathbf{x}_{2,2} \\ \vdots & \vdots & \vdots \\ 1 & \mathbf{x}_{1,n} & \mathbf{x}_{2,n} \end{bmatrix} \begin{bmatrix} \beta_0 \\ \beta_1 \\ \beta_2 \end{bmatrix} + \begin{bmatrix} e_1 \\ e_2 \\ \vdots \\ e_n \end{bmatrix}$$

# Defining models with $\mathbf{X}$

## Regression

What would  $\mathbf{X}$  look like for model with an intercept and linear increase over time  $t$ ?

$$y_t = \beta_0 + \beta_1 t + e_t$$

⇓?

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{e}$$

# Defining models with $X$

Regression

$$y = X\beta + e$$

$\Downarrow$

$$\begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{bmatrix} = \begin{bmatrix} 1 & ? \\ 1 & ? \\ \vdots & \vdots \\ 1 & ? \end{bmatrix} \begin{bmatrix} \beta_0 \\ \beta_1 \end{bmatrix} + \begin{bmatrix} e_1 \\ e_2 \\ \vdots \\ e_n \end{bmatrix}$$

# Defining models with $X$

Regression

$$y = X\beta + e$$

$\Downarrow$

$$\begin{bmatrix} y_1 \\ y_2 \\ y_3 \\ \vdots \\ y_n \end{bmatrix} = \begin{bmatrix} 1 & 1 \\ 1 & 2 \\ 1 & 3 \\ \vdots & \vdots \\ 1 & n \end{bmatrix} \begin{bmatrix} \beta_0 \\ \beta_1 \end{bmatrix} + \begin{bmatrix} e_1 \\ e_2 \\ e_3 \\ \vdots \\ e_n \end{bmatrix}$$



# Defining models with $X$

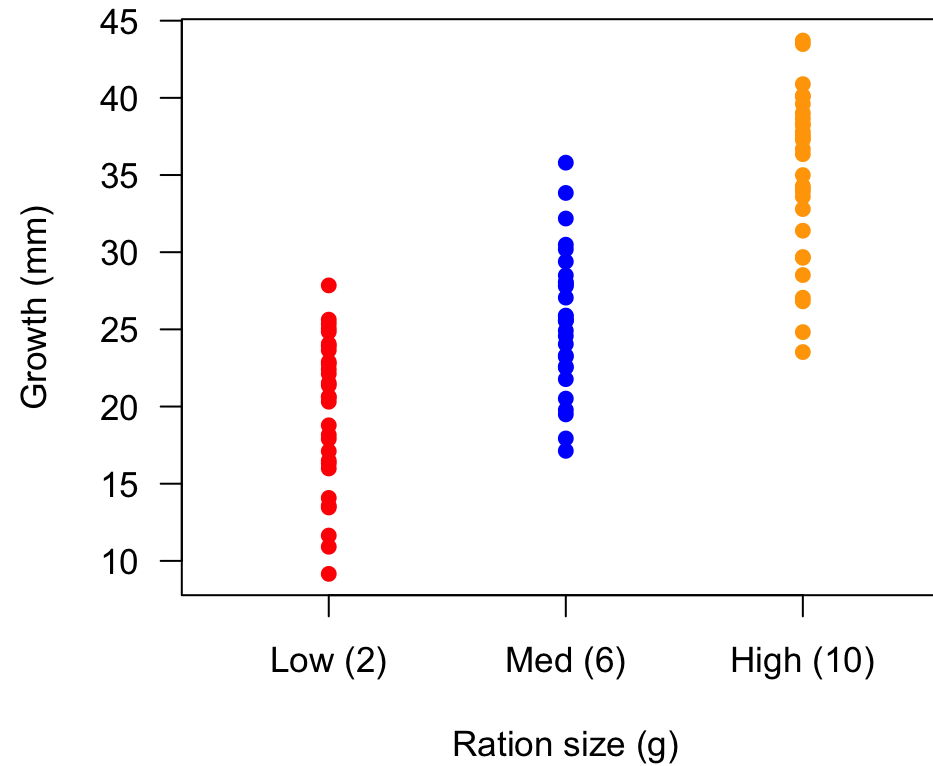
Analysis of variance (ANOVA)

ANOVA was popularized by [Ronald Fisher](#) ~100 years ago when he was studying the variance of genetic traits among commercial crops

ANOVA is used to analyze *differences among group means*

# Comparing group means

Recall our analysis of fish growth as a function of ration



# Defining models with $X$

## ANOVA

Here we want to know if the mean growth of fish varies among the 3 ration sizes

$$\bar{g}_{\text{ration}_1} \stackrel{?}{=} \bar{g}_{\text{ration}_2} \stackrel{?}{=} \bar{g}_{\text{ration}_3}$$

How would we write the model for this?

# Defining models with $X$

## ANOVA

Our model for an observation  $y_i$  is something like

$$y_i = \mu_i + e_i$$

$$\mu_i = \begin{cases} \mu_1 & \text{if fed ration 1} \\ \mu_2 & \text{if fed ration 2} \\ \mu_3 & \text{if fed ration 3} \end{cases}$$

# Defining models with $X$

## ANOVA

We can use binary 0/1 coding to represent if/then constructs

$$y_i = \mu_1 x_{1,i} + \mu_2 x_{2,i} + \mu_3 x_{3,i} + e_i$$

$x_{1,i} = 1$  if fed ration 1 and 0 otherwise

$x_{2,i} = 1$  if fed ration 2 and 0 otherwise

$x_{3,i} = 1$  if fed ration 3 and 0 otherwise

# Defining models with $X$

ANOVA

How would we specify the model matrix  $X$  for this?

# Defining models with $\mathbf{X}$

ANOVA

Let's rewrite our model as

$$y_i = \beta_1 x_{1,i} + \beta_2 x_{2,i} + \beta_3 x_{3,i} + e_i$$

↓

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{e}$$

# Defining models with $\mathbf{X}$

ANOVA

And define  $\mathbf{X}$  as

$$\mathbf{X} = \begin{bmatrix} \mathbf{x}_{1,1} & \mathbf{x}_{2,1} & \mathbf{x}_{3,1} \\ \mathbf{x}_{1,2} & \mathbf{x}_{2,2} & \mathbf{x}_{3,2} \\ \vdots & \vdots & \vdots \\ \mathbf{x}_{1,n} & \mathbf{x}_{2,n} & \mathbf{x}_{3,n} \end{bmatrix}$$



# Defining models with $X$

Let's now re-order all of the observations into their groups

$$\mathbf{y} = \begin{bmatrix} y_{1,1} \\ \vdots \\ y_{1,j_1} \\ y_{2,1} \\ \vdots \\ y_{2,j_2} \\ y_{3,1} \\ \vdots \\ y_{3,j_3} \end{bmatrix} \quad \text{with } j_1 + j_2 + j_3 = n$$

# Defining models with $\mathbf{X}$

We can then define  $\mathbf{X}$  and  $\boldsymbol{\beta}$  as

$$\mathbf{X} = \begin{bmatrix} 1 & 0 & 0 \\ \vdots & \vdots & \vdots \\ 1 & 0 & 0 \\ \hline 0 & 1 & 0 \\ \vdots & \vdots & \vdots \\ 0 & 1 & 0 \\ \hline 0 & 0 & 1 \\ \vdots & \vdots & \vdots \\ 0 & 0 & 1 \end{bmatrix} \quad \boldsymbol{\beta} = \begin{bmatrix} \beta_1 \\ \beta_2 \\ \beta_3 \end{bmatrix}$$

# Defining models with $X$

## ANOVA

Here are the mean growth rates of our 3 groups of fish

$$\bar{y}_{j=1} = \beta_1 = 19.6$$

$$\bar{y}_{j=2} = \beta_2 = 25.6$$

$$\bar{y}_{j=3} = \beta_3 = 35$$

# Defining models with $X$

## ANOVA

And here are the results of our ANOVA model

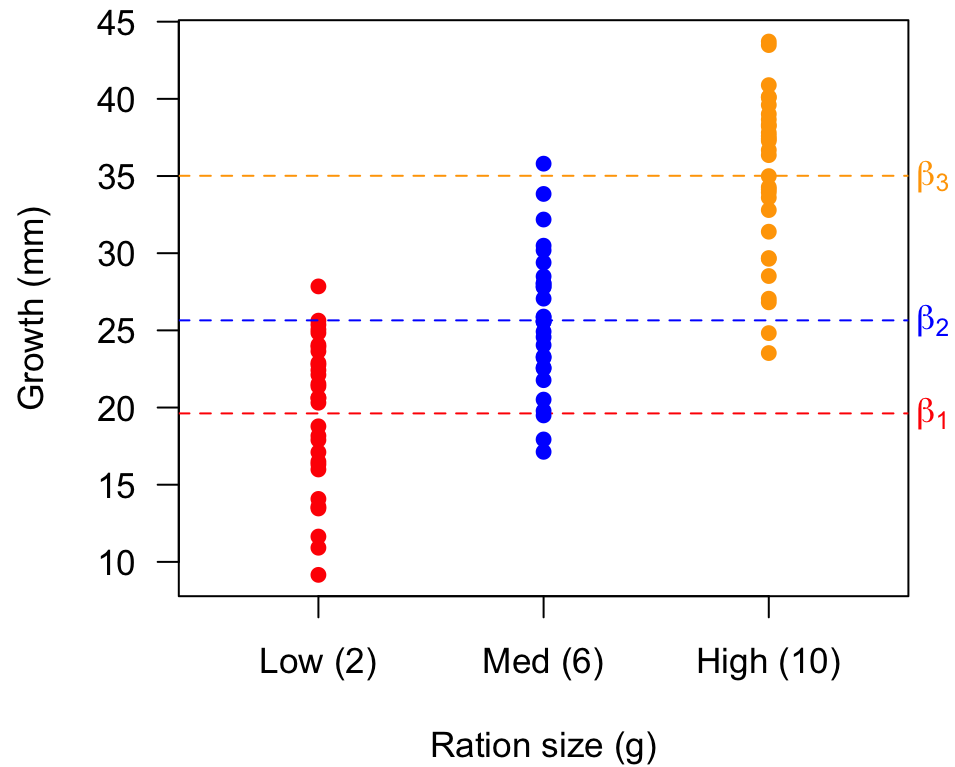
```
## fit ANOVA w/ `- 1` to remove intercept  
m1 <- lm(yy ~ ration - 1)  
coef(m1)
```

```
## ration_1 ration_2 ration_3  
## 19.62001 25.64846 35.01523
```

This confirms that we have fit a model of means

# Defining models with $X$

ANOVA



# Defining models with $X$

## ANOVA

Suppose we wanted to reframe our model to instead include the effect of ration relative to the overall mean growth rate ( $\mu$ )

$$y_i = \mu + \beta_1 x_{1,i} + \beta_2 x_{2,i} + \beta_3 x_{3,i} + e_i$$

and calculate the groups means as

$$\bar{y}_{j=1} = \mu + \beta_1$$

$$\bar{y}_{j=2} = \mu + \beta_2$$

$$\bar{y}_{j=3} = \mu + \beta_3$$

# Defining models with $\mathbf{X}$

We would then define  $\mathbf{X}$  and  $\boldsymbol{\beta}$  as

$$\mathbf{X} = \begin{bmatrix} 1 & 1 & 0 & 0 \\ \vdots & \vdots & \vdots & \vdots \\ 1 & 1 & 0 & 0 \\ \hline 1 & 0 & 1 & 0 \\ \vdots & \vdots & \vdots & \vdots \\ 1 & 0 & 1 & 0 \\ \hline 1 & 0 & 0 & 1 \\ \vdots & \vdots & \vdots & \vdots \\ 1 & 0 & 0 & 1 \end{bmatrix} \quad \boldsymbol{\beta} = \begin{bmatrix} \mu \\ \beta_1 \\ \beta_2 \\ \beta_3 \end{bmatrix}$$

# Defining models with $X$

## ANOVA

And here are the results of our ANOVA model

```
## design matrix
X <- cbind(rep(1,nn*pp), ration)
## fit ANOVA w/ ` - 1 ` to remove intercept
m2 <- lm(yy ~ X - 1)
coef(m2)
```

```
##           X           X_1           X_2           X_3
## 35.015235 -15.395221  -9.366774           NA
```

Wait-what happened here?!



# Defining models with $\mathbf{X}$

Can you spot the problem in our design matrix?

$$\mathbf{X} = \begin{bmatrix} 1 & 1 & 0 & 0 \\ \vdots & \vdots & \vdots & \vdots \\ 1 & 1 & 0 & 0 \\ \hline 1 & 0 & 1 & 0 \\ \vdots & \vdots & \vdots & \vdots \\ 1 & 0 & 1 & 0 \\ \hline 1 & 0 & 0 & 1 \\ \vdots & \vdots & \vdots & \vdots \\ 1 & 0 & 0 & 1 \end{bmatrix}$$

# Defining models with $X$

## ANOVA

```
## solve for beta by hand  
beta <- solve(t(X) %*% X) %*% t(X) %*% yy  
  
## Error in solve.default(t(X) %*% X) :  
## system is computationally singular: reciprocal condition number
```

# Defining models with $\mathbf{X}$

$\mathbf{X}$  is not *full rank* ( $\mathbf{X}_{(.1)} = \mathbf{X}_{(.2)} + \mathbf{X}_{(.3)} + \mathbf{X}_{(.4)}$ )

$$\mathbf{X} = \begin{bmatrix} 1 & 1 & 0 & 0 \\ \vdots & \vdots & \vdots & \vdots \\ 1 & 1 & 0 & 0 \\ \hline 1 & 0 & 1 & 0 \\ \vdots & \vdots & \vdots & \vdots \\ 1 & 0 & 1 & 0 \\ \hline 1 & 0 & 0 & 1 \\ \vdots & \vdots & \vdots & \vdots \\ 1 & 0 & 0 & 1 \end{bmatrix}$$

# Defining models with $X$

## ANOVA

Let's think about our model again

$$y_i = \mu + \beta_1 x_{1,i} + \beta_2 x_{2,i} + \beta_3 x_{3,i} + e_i$$

where we want the group means to be

$$\bar{y}_{j=1} = \mu + \beta_1$$

$$\bar{y}_{j=2} = \mu + \beta_2$$

$$\bar{y}_{j=3} = \mu + \beta_3$$

# Defining models with $X$

ANOVA

Consider the overall mean of  $\mathbf{y}$  in terms of the group means

$$\bar{y} = \frac{\bar{y}_{j=1} + \bar{y}_{j=2} + \bar{y}_{j=3}}{3}$$

# Defining models with $X$

## ANOVA

Consider the overall mean of  $\mathbf{y}$  in terms of the group means

$$\bar{y} = \frac{\bar{y}_{j=1} + \bar{y}_{j=2} + \bar{y}_{j=3}}{3}$$
$$\Downarrow$$
$$\mu = \frac{(\mu + \beta_1) + (\mu + \beta_2) + (\mu + \beta_3)}{3}$$
$$\Downarrow$$
$$\beta_1 + \beta_2 + \beta_3 = 0$$

# Defining models with $X$

ANOVA

Now we can rewrite our model as

$$y_i = \mu + \beta_1 x_{1,i} + \beta_2 x_{2,i} + (-\beta_1 + -\beta_2)x_{3,i} + e_i$$

and calculate the group means as

$$\bar{y}_{j=1} = \mu + \beta_1$$

$$\bar{y}_{j=2} = \mu + \beta_2$$

$$\bar{y}_{j=3} = \mu - (\beta_1 + \beta_2)$$

# Defining models with $\mathbf{X}$

We would then define  $\mathbf{X}$  and  $\boldsymbol{\beta}$  as

$$\mathbf{X} = \begin{bmatrix} 1 & 1 & 0 \\ \vdots & \vdots & \vdots \\ 1 & 1 & 0 \\ \hline 1 & 0 & 1 \\ \vdots & \vdots & \vdots \\ 1 & 0 & 1 \\ \hline 1 & -1 & -1 \\ \vdots & \vdots & \vdots \\ 1 & -1 & -1 \end{bmatrix} \quad \boldsymbol{\beta} = \begin{bmatrix} \mu \\ \beta_1 \\ \beta_2 \end{bmatrix}$$



# Defining models with $X$

## ANOVA

```
## empty design matrix
XX <- matrix(NA, nn*pp, pp)
## for mu
XX[i1,] <- matrix(c(1, 1, 0), nn, pp, byrow = TRUE)
## for beta_1
XX[i2,] <- matrix(c(1, 0, 1), nn, pp, byrow = TRUE)
## for beta_2
XX[i3,] <- matrix(c(1, -1, -1), nn, pp, byrow = TRUE)
## fit model & get parameters
Bvec <- coef(lm(yy ~ XX - 1))
names(Bvec) <- c("mu", "beta_1", "beta_2")
Bvec
```

```
##          mu    beta_1    beta_2
## 26.761236 -7.141222 -1.112776
```

# Defining models with X

## ANOVA

```
## mean of ration 1  
Bvec[ "mu" ] + Bvec[ "beta_1" ]  
## mean of ration 2  
Bvec[ "mu" ] + Bvec[ "beta_2" ]  
## mean of ration 3  
Bvec[ "mu" ] - ( Bvec[ "beta_1" ] + Bvec[ "beta_2" ] )
```

```
##      mu  
## 19.62001  
##      mu  
## 25.64846  
##      mu  
## 35.01523
```

# Defining models with $X$

## ANOVA

We could also fit our grand mean model after some simple algebra

$$y_i = \mu + \beta_1 x_{1,i} + \beta_2 x_{2,i} + \beta_3 x_{3,i} + e_i$$

↓

$$y_i - \mu = \beta_1 x_{1,i} + \beta_2 x_{2,i} + \beta_3 x_{3,i} + e_i$$

↓

$$y_i - \bar{y} = \beta_1 x_{1,i} + \beta_2 x_{2,i} + \beta_3 x_{3,i} + e_i$$

# Defining models with $X$

## ANOVA

```
## fit anova with implicit grand mean  
m2 <- lm((yy - mean(yy)) ~ ration - 1)  
coef(m2)
```

```
##  ration_1  ration_2  ration_3  
## -7.141222 -1.112776  8.253998
```

# Defining models with $X$

## ANOVA

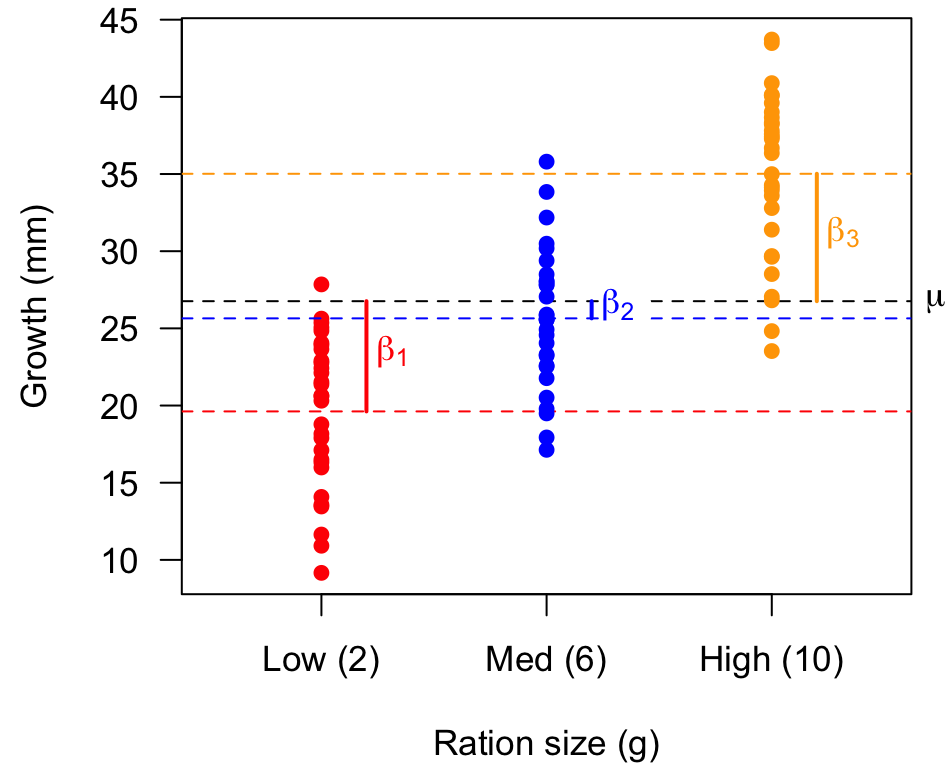
```
## do we recover our means?  
coef(m2) + mean(yy)
```

```
## ration_1 ration_2 ration_3  
## 19.62001 25.64846 35.01523
```

```
coef(m1)
```

```
## ration_1 ration_2 ration_3  
## 19.62001 25.64846 35.01523
```

# Comparing group means



# Defining models with $X$

## ANOVA

What if we wanted to treat one group as a control or reference (eg, our low ration) and estimate the other effects relative to it?

$$y_i = \beta_1 x_{1,i} + (\beta_1 + \beta_2) x_{2,i} + (\beta_1 + \beta_3) x_{3,i} + e_i$$

such that

$$\bar{y}_{j=1} = \beta_1$$

$$\bar{y}_{j=2} = \beta_1 + \beta_2$$

$$\bar{y}_{j=3} = \beta_1 + \beta_3$$

# Defining models with $\mathbf{X}$

We would define  $\mathbf{X}$  and  $\boldsymbol{\beta}$  as

$$\mathbf{X} = \begin{bmatrix} 1 & 0 & 0 \\ \vdots & \vdots & \vdots \\ 1 & 0 & 0 \\ \hline 1 & 1 & 0 \\ \vdots & \vdots & \vdots \\ 1 & 1 & 0 \\ \hline 1 & 0 & 1 \\ \vdots & \vdots & \vdots \\ 1 & 0 & 1 \end{bmatrix} \quad \boldsymbol{\beta} = \begin{bmatrix} \beta_1 \\ \beta_2 \\ \beta_3 \end{bmatrix}$$



# Defining models with $X$

## ANOVA

```
## empty design matrix
XX <- matrix(NA, nn*pp, pp)
## for beta_1
XX[i1,] <- matrix(c(1, 0, 0), nn, pp, byrow = TRUE)
## for beta_1 + beta_2
XX[i2,] <- matrix(c(1, 1, 0), nn, pp, byrow = TRUE)
## for beta_1 + beta_3
XX[i3,] <- matrix(c(1, 0, 1), nn, pp, byrow = TRUE)
## fit anova with implicit grand mean
Bvec <- coef(lm(yy ~ XX - 1))
names(Bvec) <- c("beta_1", "beta_2", "beta_3")
Bvec
```

```
##      beta_1      beta_2      beta_3
## 19.620014   6.028446  15.395221
```

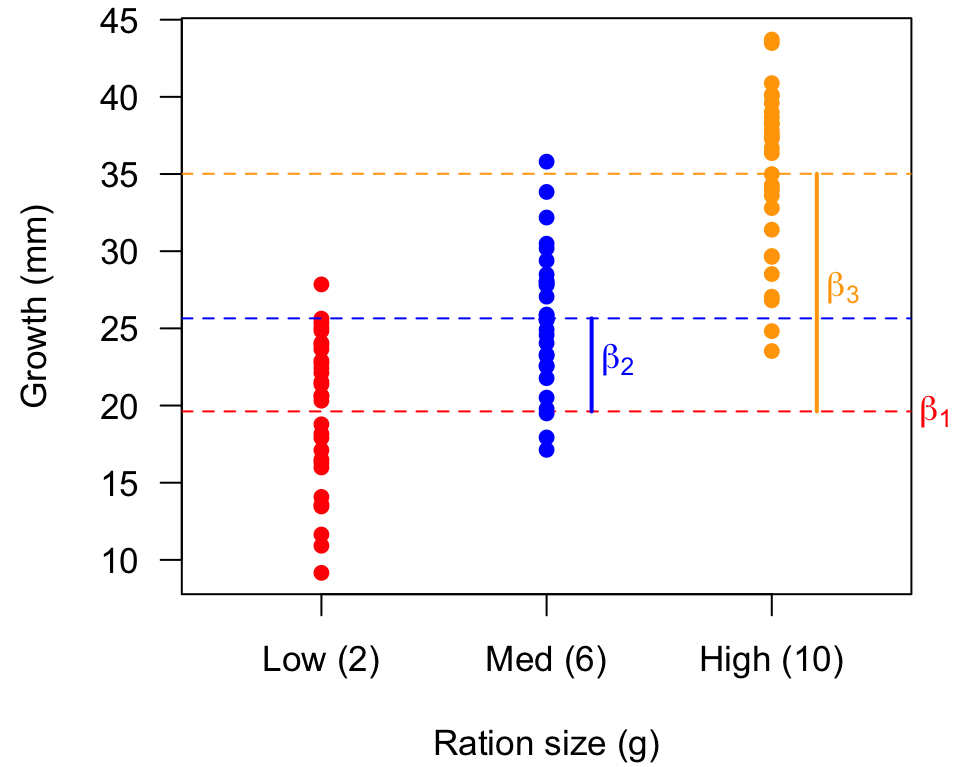
# Defining models with X

## ANOVA

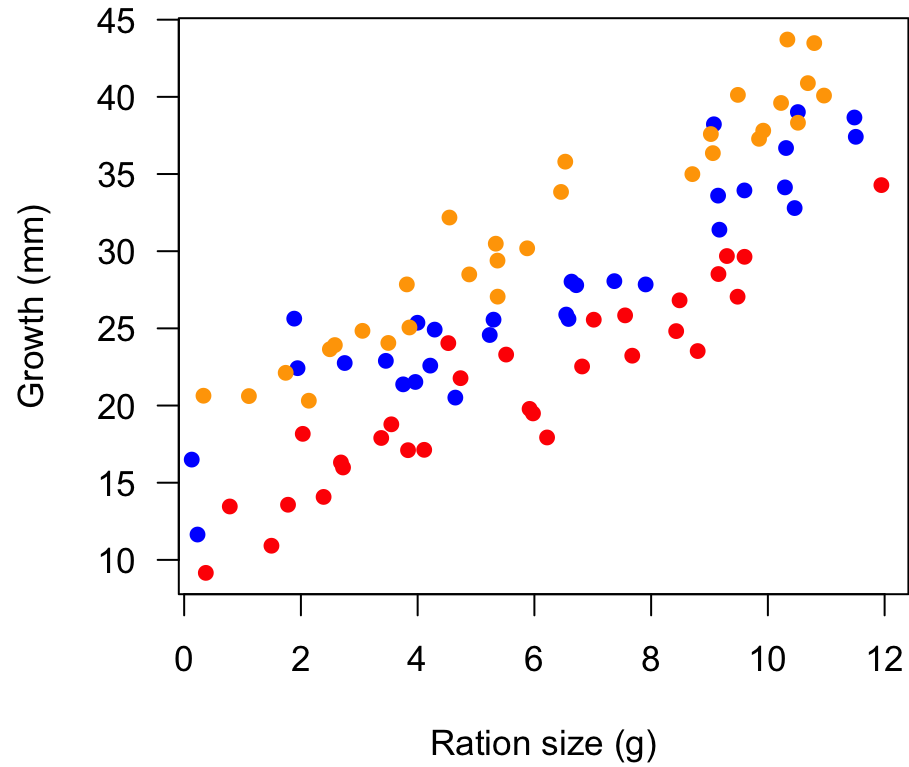
```
## mean of ration 1  
Bvec["beta_1"]  
## mean of ration 2  
Bvec["beta_1"] + Bvec["beta_2"]  
## mean of ration 3  
Bvec["beta_1"] + Bvec["beta_3"]
```

```
## beta_1  
## 19.62001  
## beta_1  
## 25.64846  
## beta_1  
## 35.01523
```

# Comparing group means



# Analysis of covariance (ANCOVA)



# Analysis of covariance (ANCOVA)

Here is our model with the categorical effect of lineage & the continuous effect of ration

$$\text{growth}_i = \alpha + \beta_{1,\text{lineage}} + \beta_2 \text{ration}_i + \epsilon_i$$

# Analysis of covariance (ANCOVA)

Dropping the global intercept & writing out the lineage effects yields

$$\text{growth}_i = \underbrace{\beta_1 x_{1,i} + \beta_2 x_{2,i} + \beta_3 x_{3,i}}_{\text{lineage}} + \underbrace{\beta_4 x_{4,i}}_{\text{ration}} + e_i$$

# Defining models with $\mathbf{X}$

We would then define  $\mathbf{X}$  and  $\boldsymbol{\beta}$  as

$$\mathbf{X} = \left[ \begin{array}{ccc|c} 1 & 0 & 0 & r_1 \\ \vdots & \vdots & \vdots & \vdots \\ 1 & 0 & 0 & r_{j_1} \\ \hline 0 & 1 & 0 & r_{j_1+1} \\ \vdots & \vdots & \vdots & \vdots \\ 0 & 1 & 0 & r_{j_2+j_2} \\ \hline 0 & 0 & 1 & r_{j_1+j_2+1} \\ \vdots & \vdots & \vdots & \vdots \\ 0 & 0 & 1 & r_n \end{array} \right] \quad \boldsymbol{\beta} = \begin{bmatrix} \beta_1 \\ \beta_2 \\ \beta_3 \\ \beta_4 \end{bmatrix}$$

# Analysis of covariance (ANCOVA)

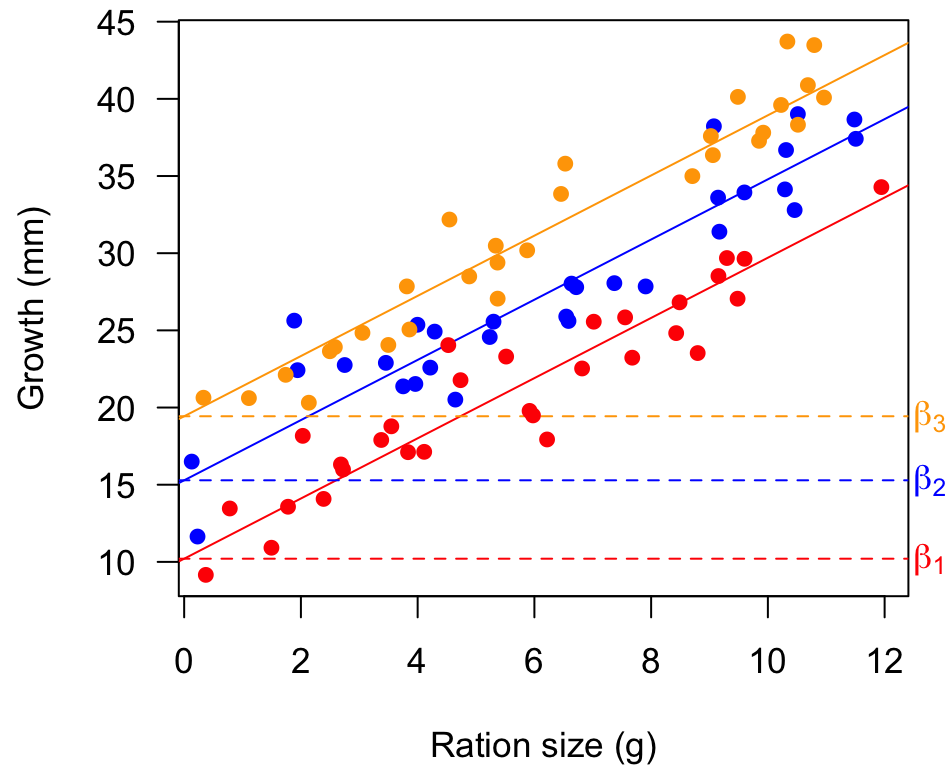
```
## create design matrix
XX <- cbind(L1 = rep(c(1,0,0), ea = nn), # effect of lineage 1
            L2 = rep(c(0,1,0), ea = nn), # effect of lineage 2
            L3 = rep(c(0,0,1), ea = nn), # effect of lineage 3
            RA = x_cov)                  # effect of ration

## fit model
Bvec <- coef(lm(yy ~ XX - 1))
names(Bvec) <- c("beta_1", "beta_2", "beta_3", "beta_4")
Bvec
```

```
##      beta_1      beta_2      beta_3      beta_4
## 10.205959 15.286507 19.435551  1.950062
```



# Analysis of covariance (ANCOVA)



# Design matrices with `model.matrix()`

We have been building our design matrices by hand, but we could instead use

`model.matrix()` with `factor()`

# Design matrices with `model.matrix()`

`factor(x)` tells R to treat `x` as categorical

```
## 2 groups with 2 obs each  
groups <- factor(c(1, 1, 2, 2))  
## inspect them  
groups
```

```
## [1] 1 1 2 2  
## Levels: 1 2
```

# Design matrices with `model.matrix()`

`model.matrix(~ x)` uses a right-hand side formula `~ x`

```
## create design matrix from `groups`  
model.matrix(~ groups)
```

```
##      (Intercept) groups2  
## 1           1      0  
## 2           1      0  
## 3           1      1  
## 4           1      1  
## attr(,"assign")  
## [1] 0 1  
## attr(,"contrasts")  
## attr(,"contrasts")$groups  
## [1] "contr.treatment"
```

# Design matrices with `model.matrix()`

What if we don't use `factor()`?

```
## 2 groups with 2 obs each
groups <- c(1, 1, 2, 2)
## create design matrix from `groups`
model.matrix(~ groups)
```

```
##      (Intercept) groups
## 1             1      1
## 2             1      1
## 3             1      2
## 4             1      2
## attr(,"assign")
## [1] 0 1
```

# Design matrices with `model.matrix()`

You can drop the intercept term with `- 1`

```
## 2 groups with 2 obs each
groups <- factor(c(1, 1, 2, 2))
## create design matrix from `groups`
model.matrix(~ groups - 1)
```

```
##   groups1 groups2
## 1      1      0
## 2      1      0
## 3      0      1
## 4      0      1
## attr(,"assign")
## [1] 1 1
## attr(,"contrasts")
## attr(,"contrasts")$groups
## [1] "contr.treatment"
```

# Design matrices with `model.matrix()`

The names/categories are irrelevant for `factor()`

```
## 2 groups with 2 obs each
groups <- factor(c("ref", "ref", "exp", "exp"))
## create design matrix from `groups`
model.matrix(~ groups)
```

```
##      (Intercept) groupsref
## 1             1             1
## 2             1             1
## 3             1             0
## 4             1             0
## attr(,"assign")
## [1] 0 1
## attr(,"contrasts")
## attr(,"contrasts")$groups
## [1] "contr.treatment"
```

# Design matrices with `model.matrix()`

R assigns factors in alphabetical order; the *reference* is first

```
## 2 groups with 2 obs each
groups <- factor(c("ref", "ref", "exp", "exp"))
## create design matrix from `groups`
model.matrix(~ groups)
```

```
##      (Intercept) groupsref
## 1             1             1
## 2             1             1
## 3             1             0
## 4             1             0
## attr(,"assign")
## [1] 0 1
## attr(,"contrasts")
## attr(,"contrasts")$groups
## [1] "contr.treatment"
```



# Design matrices with `model.matrix()`

We can change the reference case with `relevel()`

```
## 2 groups with 2 obs each
groups <- relevel(groups, "ref")
## create design matrix from `groups`
model.matrix(~ groups)
```

```
##      (Intercept) groupsexp
## 1             1          0
## 2             1          0
## 3             1          1
## 4             1          1
## attr(,"assign")
## [1] 0 1
## attr(,"contrasts")
## attr(,"contrasts")$groups
## [1] "contr.treatment"
```

# Design matrices with `model.matrix()`

We can add multiple factors with +

```
diet <- factor(c(1, 1, 2, 2))
sex <- factor(c("f", "m", "f", "m"))
model.matrix(~ diet + sex)
```

```
## (Intercept) diet2 sexm
## 1          1      0    0
## 2          1      0    1
## 3          1      1    0
## 4          1      1    1
## attr(,"assign")
## [1] 0 1 2
## attr(,"contrasts")
## attr(,"contrasts")$diet
## [1] "contr.treatment"
##
## attr(,"contrasts")$sex
## [1] "contr.treatment"
```

# Design matrices with `model.matrix()`

You can also extract the design matrix from a fitted model

```
## ANCOVA model from above
mod_fit <- lm(yy ~ XX - 1)
## get design matrix
mm <- model.matrix(mod_fit)
head(mm)
```

```
##      XXL1 XXL2 XXL3      XXRA
## 1      1     0     0 11.944444
## 2      1     0     0  3.835147
## 3      1     0     0  3.376075
## 4      1     0     0  4.112188
## 5      1     0     0  2.721664
## 6      1     0     0  1.779256
```